
Original Article

Leading or lagging indicators of risk? The informational content of extra-financial performance scores

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ABSTRACT This study investigates the informational content of extra-financial agency scoring by examining the relationship between firm beta and extra-financial performance score upgrades and downgrades. Specifically, we study the variations in the extra-financial



score of 266 Canadian corporations between 2007 and 2012 with a conditional model. We find no evidence that changes in firm beta precede changes in extra-financial scores. Rather, our results suggest that a firm's systematic risk increases following a downgrade of its extra-financial performance. In terms of score upgrades, the overall effect is not significant. However, score upgrades for firms with already-high scores predict higher systematic risk, while score upgrades for firms with low scores predict lower systematic risk. These results suggest that extra-financial scores are informational and can be useful to portfolio managers, notably for their risk management strategies.

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INTRODUCTION

Firm spending in corporate social responsibility (CSR) activities has substantially increased in recent years. According to Hong *et al* (2012), CSR spending amounts to hundreds of millions of dollars annually. One reason for this renewed interest for socially responsible investments (SRI) is that investors, including institutional investors, are becoming more concerned with the extra-financial consequences of corporate decisions. Firms that neglect this aspect may therefore face greater financial risk due to possible actions by regulators and activists that will affect their profitability (e.g., Baron and Diermeier, 2007; Lyon and Maxwell, 2011).

Growing enthusiasm for SRI has led to a surge in rating agencies that specialize in social and environmental rating and scoring (e.g., MSCI ESG STATS¹ in the US; EIRIS² in the UK; Thomson Reuter's ASSET4³ and Sustainalytics⁴ which operate globally). In addition to their main role of providing investors with information on CSR strategies, some agencies also publish extra-financial performance scores.⁵ Our study examines whether variations in extra-financial scores affect financial markets, similar to what changes in credit ratings do (e.g., Jorion and Zhang, 2007; Holthausen and Leftwic, 1986; Weinstein, 1977). Specifically, we investigate

the information content of extra-financial scores by examining their relationships to firms' systematic risk (beta). We address the following two questions: (1) is beta related to extra-financial performance scores and, (2) if so, do extra-financial performance score changes (upgrades and downgrades) lead or lag indicators of beta? Knowing whether firm beta varies before or following the extra-financial performance changes is an important practical question. Specifically, if extra-financial performance changes predict systematic risk changes, then extra-financial rating agencies' scores can be an excellent risk management tool, particularly for institutional investors. In theory, extra-financial performance scores can be leading indicators of corporate systematic risk if extra-financial rating agencies are able, through their analysis of a company's environment, social or governance (ESG) criteria, to predict future losses or risk events such as operational or reputational losses. If, however, rating agencies are mostly reacting to corporate events that are related to extra-financial performance, then extra-financial performance score changes will lag beta variations. In this case, extra-financial agencies' scores would be less useful as predicting tools.

Our study contributes to the literature in several ways. Firstly, while previous studies



investigate the impact of extra-financial performance score *levels* on financial risk (e.g., Kim *et al*, 2014; Bouslah *et al*, 2013; Oikonomou *et al*, 2012), our research focuses on the impacts of extra-financial performance score *changes* on systematic risk and by distinguishing between the effects of score upgrades and downgrades. From a risk manager's point of view, score changes are fundamentally different from score levels, as they are related to *new* information about a firm's risk. For instance, a firm can experience a score downgrade and still maintain a high score, or experience a positive change in its extra-financial score and still have a low score.

Secondly, while most studies agree on the existence of a relationship between extra-financial performance and financial risk, as evidenced by Orlitzky and Benjamin's (2001) meta-analysis, the question of the direction of the relationship between the two variables is not yet settled. Some authors argue that systematic risk is a determinant of CSR, as managers in lower-risk companies have access to more stable cash flows, allowing them to improve their extra-financial performance (Hasseldine *et al*, 2005; Roberts, 1992; McGuire *et al*, 1988). Furthermore, Krüger (2015) shows that the occurrence of firm-specific events, related to environmental, social or governance risks, has an important influence on KLD's scorings. Specifically, he shows that KLD's scores are updated to account for information on events that have already occurred. Others believe that an improvement in a firm's extra-financial performance is likely to be rewarded by the market in terms of improved risk perception, and thus by a lower beta (e.g., Oikonomou *et al*, 2012; Salama *et al*, 2011; Sharfman and Fernando, 2008). Further, the previous studies do not address the basic question: whether a firm's financial risk is low (high) because of its high (low) CSR or whether its CSR is high (low) because of its low (high) financial risk? The mere observation of a negative correlation

between some annual CSR measures and financial risk is consistent with at least two different interpretations: either more responsible firms tend to be less risky or, alternatively, less-risky firms tend to channel more resources into projects that increase their CSR. Our extra-financial performance data allow us to test whether beta variations occur before or following changes in firms' extra-financial performance.

Thirdly, our study addresses the omnipresent issue of over-investment and managerial opportunism (e.g., McWilliams and Siegel, 2001; Preston and O'Bannon, 1997), which suggests that, under certain circumstances, extra-financial performance can be a potential source of risk, for instance, because of overinvestment (McWilliams and Siegel, 2001). To do so, we examine whether extra-financial performance score upgrades for firms with already high scores predict higher systematic risk. Finally, in order to reflect the qualitative differences across the dimensions of extra-financial performance, we separately analyze each of the three components of extra-financial performance (i.e., environment, social and governance) as well as in an aggregate measure of performance.

Our results show no significant evidence that extra-financial score changes lag beta variations. Rather, we observe that systematic risk increases following extra-financial score downgrades. The overall predictive power of score upgrades is not significant. Therefore, extra-financial performance scores are not simply reacting to market information but are, particularly downgrades, leading indicators of firm systematic risk variations. Further, we show that extra-financial score upgrades for firms with already-high scores predict higher systematic risk while they predict lower systematic risk for firms with low extra-financial scores. This result suggests that, for firms with already-high extra-financial scores, further improvements can be counterproductive and lead to an increase in long-term risk, possibly because of costs that



investors feel are too high and inopportune. By contrast, systematic risk decreases when firms with low extra-financial scores make an effort to improve their social image. Regarding score downgrades, as for our general results, we find that they predict higher systematic risk for firms with low extra-financial scores.

The rest of the study is structured as follows. “[Prior Research on the Impact of Extra-Financial Performance on Financial Risk](#)” section presents a summary of the literature on the impact of extra-financial performance on shareholder wealth (return and risk). “[Theoretical Framework and Research Hypotheses](#)” section presents the theoretical framework and research hypotheses. “[Data and Methodology](#)” section describes the data and the methodology used in order to test our hypotheses. “[Empirical Results and Discussion](#)” section presents and discusses our empirical results, and finally, “[Conclusion](#)” section concludes the paper.

PRIOR RESEARCH ON THE IMPACT OF EXTRA-FINANCIAL PERFORMANCE ON FINANCIAL RISK

Unlike the abundant literature on the impact of extra-financial performance on firm financial performance, there are few studies that examine the relationship between financial risk and extra-financial performance. These few studies analyze different measures of financial risk, such as variance and its components (idiosyncratic risk, and systematic risk). Some studies suggest that extra-financial performance affects only idiosyncratic risk because extra-financial performance is firm specific. For example, using data between 1995 and 1999 from the “Canadian Social Investment database”, Boutin-Dufresne and Savaria (2004) find a negative relationship between CSR and firm idiosyncratic risk. This observation is confirmed by Lee and Faff (2009), who study the impact of CSR on financial risk for firms

listed in the Dow Jones Sustainability Index. The authors demonstrate that socially responsible firms are less risky than their socially irresponsible counterparts. Using Fortune’s MAC data between 2002 and 2003 as a measure of CSR, Luo and Bhattacharya (2009) show that CSR decreases a firm’s idiosyncratic risk and provides insurance against the volatility of the firm’s future cash-flows. Mishra and Modi (2012) and Bouslah *et al* (2013) confirm this result by using KLD data as a principal proxy for extra-financial performance. Mishra and Modi (2012) observe that CSR has a significant effect on idiosyncratic risk over the period spanning from 2000 to 2009, with positive CSR scores reducing risk and negative CSR scores increasing it. Bouslah *et al* (2013) focus their analysis on the individual components of extra-financial performance. They find that financial risk (measured by idiosyncratic risk as well as stock return volatility) is negatively related to two CSR components, namely employee relations and human rights, while other CSR components do not affect financial risk.

Several studies argue that investigating the effects of CSR on systematic risk is more relevant because, in the absence of market imperfections, only systematic risk is priced; idiosyncratic risk can be eliminated through diversification. McGuire *et al* (1988) find that CSR, proxied by firm ranking in Fortune’s list of America’s most admired companies (MAC), is negatively related to market risk loadings over the period 1983–1985. In this study, beta is a lagged independent variable, so that low financial risk is theorised to create the planning certainty that facilitates investment in CSR. Unlike McGuire *et al* (1988) and Luo and Bhattacharya (2009) consider systematic risk as the dependent variable and simultaneously control for small cap and book-to-value effects in their systematic and idiosyncratic risk estimations. The authors conclude that a firm’s extra-financial performance, as evaluated by Fortune magazine, is negatively correlated with beta measures. Salama *et al* (2011)



address this issue in the UK by examining firm activity from 1994 to 2006. They also consider systematic risk as the dependent variable and predominantly focus on environmental responsibility. Their results do not stray too far from those related to CSR in the American context and attest that the environmental performance of UK firms is inversely related to systematic risk. Jo and Na (2012) find that a firm's overall CSR engagement alleviates not only total risk, but also systematic risk and sensitivities to market fluctuations, particularly for controversial industries in the US. Also in the US context, Oikonomou *et al* (2012) present a longitudinal study that analyzes the relationship between corporate extra-financial performance and systematic risk between 1992 and 2009 using the KLD database and find that CSR is negatively related to systematic risk. The authors find a negative (positive) relation between systematic risk and a measure of aggregate extra-financial strengths (concerns). However, they also find that only community, employment, and environmental concerns are significantly and positively related to systematic risk. The authors also note that the impact of ESG criteria on firm risk varies according to the economic context measured with market volatility. Lastly, Kim *et al* (2014) show that if socially responsible firms commit to a high standard of transparency they would have lower crash risk. However, if managers engage in CSR to cover up bad news and divert shareholder scrutiny, CSR would be associated with higher crash risk.

THEORETICAL FRAMEWORK AND RESEARCH HYPOTHESES

There are two major theoretical arguments that link corporate extra-financial performance to financial risk. The first posits that a high extra-financial performance brings about extra operating costs and potential

sacrifices and, hence, puts firms with high extra-financial performance scores into a risk disadvantage. The second argument, based on the stakeholder theory, contends that although increasing its extra-financial performance can be costly for a firm, it can reduce other costs and/or improve revenues and thereby decrease financial risk.

Stakeholder theory

Stakeholder theory states that every modern firm has explicit and implicit relationships with a variety of stakeholders who have the power to determine its success or failure (e.g., Jones, 1995; Wijnberg, 2000).

The advantages of adopting a CSR approach that takes into account stakeholders' interest are multiple and go with the principles of a risk management system main objective of which is to prevent or avoid the disruption, loss or damage to business operations. For example, the fact that all stakeholders (including shareholders) feel more involved in the decision-making process reduces information asymmetry (see, for e.g., Waddock and Graves, 1997) and uncertainty about future cash flows (e.g., McGuire *et al*, 1988). Sharfman and Fernando (2008) argue that risk management of social or environmental issues is theoretically synonymous with strategic risk management because it reduces potential risks (e.g., accidents, labor disputes, consumer boycotts, damage to brand image, and reputation), lowers favorable investor recognition and, consequently, reduces the number of potential claimants on a firm's cash flows (e.g., potential fines, compliance cost, etc.). Sharfman and Fernando (2008) conclude that, when potential litigations are reduced, cash flows are more stable, and a firm's resources can be dedicated to strategic decisions and investments that contribute to reducing the financial risk perceived by the market (i.e., systematic risk).

However, the failure of firms to meet the claims of implicit stakeholders can result in



costly explicit claims (e.g., lawsuits, regulatory intervention etc.) to force their hands. Investors can anticipate this situation and consider investment in these firms as risky (e.g., Stern, 2006; Porter and Kramer, 2006). Assuming that stakeholder claims are of similar nature across all firms, these collective claims might lead to a systematic event, such as a downturn in the economic cycle or a change in the legislative framework and can have systematic effects on all firms or common groups of firms.

Our study attempts to test the informational content of extra-financial rating agency scorings by examining the relationship between firm betas and extra-financial performance score changes. Following the above discussion, we expect that downgrades (upgrades) in extra-financial performance scores are related to increases (decreases) in firms' systematic risk. This is summarized in our first research hypothesis:

H1 Extra-financial performance score changes are related to firms' systematic risk variations.

As mentioned previously, while there seems to be consensus on the relationship between extra-financial performance and financial risk, as evidenced by Orlitzky and Benjamin (2001), the direction of the causal link between the two variables is still an empirical issue. Roberts (1992) and Hasseldine *et al* (2005) argue that systematic risk is a determinant of CSR, as managers in lower-risk companies have access to more stable cash flows, allowing them to improve their extra-financial performance. Others believe that the improvement of extra-financial performance is likely to be rewarded by the market in terms of improved risk perception, and thus by a lower beta (e.g., Oikonomou *et al*, 2012; Salama *et al*, 2011; Sharfman and Fernando, 2008). To address this dual link, we decompose *H1* into two testable subhypotheses (*H1a* and *H1b*). If extra-financial rating agencies are able, through their analysis of a company's ESG criteria, to predict future losses or risk events

(such as operational or reputational losses), then extra-financial scores should be leading indicators of systematic risk. This is highlighted in subhypothesis *H1a*:

H1a Extra-financial performance scores are leading indicators of systematic risk.

If, however, rating agencies are mostly reacting to corporate events that are related to extra-financial performance, then extra-financial performance scores will be lagging indicators of systematic risk. This is summarized in subhypothesis *H1b*:

H1b Extra-financial performance scores are lagging indicators of systematic risk.

We anticipate the relationship between extra-financial performance and financial risk to be asymmetrical. As argued by some authors (e.g., Mattingly and Berman, 2006; or Oikonomou *et al*, 2012), it is unreasonable to assume that stakeholders will react to responsible and irresponsible behaviors in opposite yet symmetrical manners. Furthermore, there is recent empirical evidence that CSR and corporate social *irresponsibility* affect a firm's bottom line to differing magnitudes (Krüger, 2015). This is summarized in the second research hypothesis:

H2 The informational content of extra-financial performance score downgrades is higher than score upgrades.

Over-investment and managerial opportunism theories

Over-investment and managerial opportunism theories (e.g., McWilliams and Siegel, 2001; Preston and O'Bannon, 1997) support a positive relationship between extra-financial performance and financial risk. According to the proponents of these theories, managers may choose to improve their firm's extra-financial performance score at the expense of shareholders by over-investing in CSR activities in order to build their own personal reputation as good social citizens



(Barnea and Rubin, 2010) or to generate support from social and environmental activists, local communities, politicians, NGOs, etc., in order to reduce the probability of their replacement in a future period (Cespa and Cestone, 2007) or even to hide bad management (Hemingway and Maclagan, 2004). This strategy, if known, would be sanctioned by a higher financial risk. Lastly, McWilliams and Siegel (2001) believe that there is an optimal level of extra-financial performance, beyond which it is less likely to shield the firm against the uncertainty and vulnerability of future cash flows.

At very high levels of extra-financial performance, the disadvantages of CSR in the context of a firm's economic purposes may outweigh its benefits, thus likely inducing more unstable future profits and less insurance-like protection against stock return risk. This is summarized in our risk-related third research hypothesis:

H3 The relationship between extra-financial performance and financial risk is stronger for firms with already-high extra-financial performance scores.

DATA AND METHODOLOGY

Data

We use corporate social ratings data from the Sustainalytics database. Sustainalytics specialises in the measurement of corporate extra-financial performance against a predetermined set of criteria, as shown in Appendix, and is principally used by institutional investors. Unlike MSCI ESG, which evaluates CSR based on seven qualitative criteria, Sustainalytics scores firms on over 100 proprietary indicators for the three ESG criteria. Furthermore, unlike MSCI ESG, which assigns positive and negative ratings (i.e., strengths and concerns), Sustainalytics' extra-financial performance scores range from 0 (worst) to 10 (best). Daily returns and macroeconomic variables used herein are

collected from the Canadian Financial Markets Research Center (CFMRC) and Bloomberg databases.

Our final sample consists of 266 publicly traded firms listed on the Toronto Stock Exchange for which there are at least two observations from January 2007 to December 2012 in the Sustainalytics database. In total, 2213 extra-financial score changes are studied, consisting of 1312 upgrades and 901 downgrades. Table 1 presents descriptive statistics on extra-financial score levels and changes for the firms in our sample. Score changes and levels are presented for the aggregate score and by dimension (environmental, social, and governance). Extra-financial performance changes appear to be asymmetric, as evidenced by the skewness coefficients reported in Panel A of Table 1. Specifically, score changes are negatively skewed for the environment and social dimensions and positively skewed for the governance dimension as well as for the aggregate score. The Jarque-Bera statistics are significant for all series and confirm that an assumption of normality is not verified. In addition, statistical tests (t tests for the mean and Wilcoxon signed-rank test for the median) do not reject the null hypotheses for both the mean and median of extra-financial performance changes at the 99 per cent level of significance.

From Panel B, which shows score change statistics for upgrades and downgrades, we note that score upgrades are more frequent than score downgrades, as evidenced by the higher number of observations in the first case.

Methodology

Both theoretical developments and empirical evidence suggest that systematic risk is not constant, but changes over time (e.g., Ferson and Schadt, 1996; Christopherson *et al*, 1998; Champagne *et al*, 2015). These changes are related to predetermined information variables. Our empirical model extends the

Table 1: Descriptive statistics on extra-financial performance scores

Extra-financial performance (score/10)	Panel A: descriptive statistics for extra-financial performance score levels and changes															
	Level					Change					Obs.					
	Mean	Std	Min	Max	Med.	Skew.	Kurt.	Mean	Std	Min		Max	Med.	Skew.	Kurt.	Jarque–Berra
Aggregate	5.467*	0.837	2.700	7.954	5.400*	-0.258	0.300	0.022	1.282	-6.708	6.780	0.023	0.065	3.010	837.13*	2213
Environment	5.387*	1.000	2.929	8.900	5.301*	0.313	0.013	0.023	1.317	-7.657	5.932	0.024	-0.016	3.012	709.34*	1876
Social	5.458*	1.479	1.000	9.586	5.510*	-0.280	-0.013	0.031	1.406	-6.277	7.602	0.030	-0.075	2.981	538.80*	1451
Governance	6.493*	1.065	3.500	9.686	6.500*	0.050	-0.629	-0.005	1.421	-6.577	6.612	-0.005	0.088	2.978	640.33*	1727

Extra-financial performance (score/10)	Panel B: descriptive statistics for extra-financial performance score upgrades and downgrades												
	Score upgrades					Score downgrades					N	# of firms	
	Mean	Std	Min	Max	Med.	Observations	Mean	Std	Min	Max			Med.
Change	0.206	0.210	0.002	2.074	0.534	1312	-0.141	0.159	-2.400	-0.002	-0.530	901	244
Aggregate	0.565	0.598	0.002	3.547	0.581	1104	-0.457	0.573	-5.031	-0.014	-0.430	772	235
Environment	0.713	0.680	0.000	5.172	0.735	930	-0.498	0.493	-3.093	-0.004	-0.696	521	221
Social	0.572	0.552	0.004	4.360	0.524	867	-0.693	0.667	-4.360	-0.088	-0.523	860	254
Governance													

Note: This table presents descriptive statistics for the extra-financial performance scores for a sample of 266 Canadian firms between 2007 and 2012. Statistics are presented for the aggregate score and for each extra-financial score dimension: Environment, Social, and Governance. Panel A presents statistics for score levels and changes, while Panel B shows statistics for score upgrades and downgrades. *N* is the number of observations used to estimate the statistics. ** indicate rejection of the null hypotheses for the mean, median, and normality at the 99 per cent significance level for statistical tests (respectively *t* test for the mean, Wilcoxon signed-rank test for the median, and Jarque–Berra test for normality).



conditional modeling approach suggested by Christopherson *et al* (1998), by adding extra-financial performance score changes.⁶ Formally, similar to the methodology used by Champagne *et al* (2015), our conditional model incorporates changes in extra-financial scores as well as publicly available economic instruments into financial performance estimation to account for the possibility of time variation in betas and abnormal performance (alpha). After integrating Fama–French’s (1993) factors, our empirical model is of the form:

$$\begin{aligned}
 R_{it} - R_{ft} &= \alpha_{it}(\delta_{Et}, Z_{n,t-1}) \\
 &+ \beta_{it}(\delta_{Et}, Z_{n,t-1})(R_{mt} - R_{ft}) \\
 &+ \beta_{1i}SMB_t + \beta_{2i}HML_t + \beta_{3i}Jan_t \\
 &+ \beta_{4i}Mon_t + \varepsilon_{it}, \quad (1)
 \end{aligned}$$

where $R_{it} - R_{ft}$ is the excess daily return of firm i on day t . R_{it} and R_{ft} respectively, designate the return for firm i and the risk-free rate (i.e., the daily yield on a 90-day maturity government bond) on day t . Market portfolio return R_{mt} is the value-weighted stock return of the S&P-TSX index. Risk factors HML_t and SMB_t respectively represent the book-to-market ratio effect and the size effect (Fama and French, 1993). Jan_t and Mon_t are binary variables that control for the January and Monday effects, respectively, and equal 1 for the control period and 0 otherwise. δ_{Et} is a dummy variable that equals 1 if day t is included in the event window (i.e., starting when we observe a change in firm i ’s extra-financial score), and 0 otherwise. Vector $Z_{n,t-1}$ with $n = 1, 2, \dots, 5$, includes the five⁷ macroeconomic information variables that condition beta, $\beta_{it}(\delta_{Et}, Z_{n,t-1})$. ε_{it} is the error term for firm i and $\varepsilon_{it} \sim N(0, \sigma)$.

Conditional beta for firm i is defined as follows:

$$\beta_{it}(\delta_{Et}, Z_{n,t-1}) = b_{0i} + b_{Ei}\delta_{Et} + \sum_{n=1}^5 b_{ni}z_{n,t-1}, \quad (2)$$

where $z_{n,t-1} = Z_{n,t-1} - E(Z_n)$ is a vector of the deviations of $Z_{n,t-1}$ from the unconditional means. b_{0i} measures the average conditional beta unrelated to score changes and macroeconomic information variables. Parameters b_{ni} (for $n = 1, \dots, 5$) measure conditional beta’s sensitivity to the five macroeconomic information variables, $Z_{n,t-1}$. b_{Ei} , represents beta variations associated with extra-financial performance score changes (upgrades or downgrades). More specifically, b_{Ei} measures the difference between beta estimated with the model that takes into account the changes in extra-financial firms scores and beta estimated without taking these changes into account. Formally, since:

$$\begin{aligned}
 \text{for } \delta_{Et} = 1, \beta_{it}(1, Z_{n,t-1}) &= b_{0i} + b_{Ei}.1 + \\
 \sum_{n=1}^5 b_{ni}z_{n,t-1} \text{ and for } \delta_{Et} = 0, \beta_{it}(0, Z_{n,t-1}) &= \\
 b_{0i} + b_{Ei}.0 + \sum_{n=1}^5 b_{ni}z_{n,t-1},
 \end{aligned}$$

we obtain:

$$\begin{aligned}
 b_{Ei} &= \beta_{it}(1, Z_{n,t-1}) - \beta_{it}(0, Z_{n,t-1}) \\
 &= (b_{0i} + b_{Ei}.1 + \sum_{n=1}^5 b_{ni}z_{n,t-1}) \\
 &\quad - (b_{0i} + b_{Ei}.0 + \sum_{n=1}^5 b_{ni}z_{n,t-1}). \quad (3)
 \end{aligned}$$

We estimate t -stats for the models using the heteroskedasticity-consistent estimation techniques of Newey and West (1987).

To test *H1a* and *H1b*, we estimate the average coefficient b_{Ei} around changes in extra-financial performance scores. More specifically, we estimate this average coefficient on analysis periods of 60 and 120 days before changes in extra-financial scores (i.e., $[-60; 0]$, $[-120; 0]$), and on analysis periods of 60, 120 and 250 days following changes in extra-financial scores (i.e., $[0; +60]$, $[0; 120]$, and $[0; +250]$).

**Table 2:** Conditional model and time-varying beta

Variable	Score upgrades				Score downgrades			
	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.
SMB	0.1824 (7.32)	0.1980 (9.66)	0.2052 (6.23)	0.1999 (8.81)	0.1799 (8.70)	0.2236 (6.72)	0.2190 (9.51)	0.1671 (8.63)
HML	0.1358 (5.85)	0.1240 (5.26)	0.1147 (4.17)	0.1260 (4.19)	0.1092 (5.27)	0.1318 (5.38)	0.1192 (5.35)	0.1399 (5.14)
Jan	-0.0001 (-1.26)	0.0001 (0.52)	-0.0002 (-1.77)	-0.0002 (-1.61)	0.0000 (0.02)	0.0000 (0.17)	-0.0001 (-0.31)	-0.0001 (-0.40)
Mon	-0.0004 (-6.36)	-0.0004 (-4.39)	-0.0004 (-4.49)	-0.0005 (-5.88)	-0.0003 (-3.29)	-0.0003 (-3.03)	-0.0004 (-2.96)	-0.0003 (-2.86)
b_{0i}	0.0003 (1.46)	0.0005 (1.85)	0.0002 (0.58)	0.0000 (-0.03)	0.0002 (0.77)	0.0003 (1.00)	0.0004 (1.01)	0.0003 (0.972)
Short-term rate	-0.0372 (-4.64)	-0.0290 (-3.01)	-0.0430 (-3.76)	-0.0329 (-3.07)	-0.0315 (-2.95)	-0.0253 (-2.14)	-0.0768 (-4.97)	-0.0082 (-2.01)
Term structure slope	0.0272 (5.43)	0.0233 (5.68)	0.0272 (2.63)	0.0264 (2.87)	0.0271 (4.20)	0.0290 (6.85)	0.0279 (4.82)	0.0233 (5.80)
Stock market return	0.0368 (4.22)	0.0385 (4.90)	0.0318 (3.11)	0.0273 (2.47)	0.0321 (3.07)	0.0340 (3.16)	0.0299 (2.98)	0.0249 (1.98)
Stock market implied volatility	0.0209 (2.42)	0.0263 (2.68)	0.0254 (2.78)	0.0264 (3.87)	0.0215 (3.20)	0.0239 (4.53)	0.0247 (2.63)	0.0251 (1.91)
Credit spread	0.0447 (10.34)	0.0551 (10.40)	0.0593 (10.40)	0.0468 (8.61)	0.0573 (10.22)	0.0732 (11.86)	0.0354 (4.79)	0.0535 (8.90)
Adj- R^2	0.228	0.236	0.238	0.229	0.231	0.237	0.237	0.224
F-stat	7092.45	5426.44	4266.14	4281.57	4649.01	4122.91	2513.37	3402.86
Obs.	1312	1104	930	867	901	772	521	860

Note: This table presents the regression results from the estimation of a different specification of model (1) in which the score change components have been removed. The coefficients for the remaining four variables: SMB, HML, Jan, and Mon are presented in the top part of the table (variables are defined in section "Data and Methodology"). The bottom part of the table shows the results for the specification of model (2) in which the score change component has been removed. The coefficients for b_{0i} , which represents the average conditional beta, and for each of the five macroeconomic variables (Z_i) (i.e., the short-term interest rates, the term structure slope, the stock market return, the stock market (implied) volatility, and the credit spread) are presented in the bottom part of the table. The analysis period is [-500; 250] days. Models are estimated on a subsample of score upgrade observations and a subsample of score downgrades and, in each case, for the four types of score dimensions [aggregate (agg.), environment (envir.), social, and governance (gov.)]. Our sample includes 266 Canadian corporations from January 2007 to December 2012. The estimated coefficients' mean values are presented, with t statistics in parentheses. Numbers in bold indicate significance at the 10 per cent level.

EMPIRICAL RESULTS AND DISCUSSION

Preliminary results

We first examine whether the conditional model framework commonly used in the US context is appropriate in this study which focuses on the Canadian context. To do so, we estimate models (1) and (2) without the term $b_{Ei}\delta_{Ei}$ over the [-500; 250] window, which corresponds to the 250 days following changes in extra-financial scores. Table 2 shows the results of these regressions. We note that our augmented conditional Fama-

French model is relevant. Specifically, the coefficients for both HML and SMB are positive and significant. Further, while the January (Jan) effect is not clear, the Monday (Mon) effect is relevant with a positive and significant coefficient. Finally, coefficients for macroeconomic information variables ($Z_{n,t-1}$) are highly significant, which indicate that systematic risk (beta) is a function of the economic context. The use of a conditional model is therefore justified to disentangle the impact of the economic context from the impact of extra-financial score changes on corporate financial risk. Because extra-financial performance and macroeconomic



factors are known to co-vary (e.g., Albuquerque *et al*, 2014; Oikonomou *et al*, 2012; Chen *et al*, 2010), the ability to separate the two effects is crucial. The fact that our sample period covers the period from 2007 to 2012, which is characterized by financial turmoil, provides an even stronger case for the use of a conditional model.

Beta variations around changes in extra-financial performance

Table 3 presents the mean values for coefficients b_{Ei} that measure beta variations around changes in firms' extra-financial performance scores [see model (2)] for both the unconditional model (estimated without the information variables, $Z_{n,t-1}$) and conditional model.⁸ Results show extra-financial score changes are negatively related to systematic risk, irrespective of the model used. Specifically, score upgrades are related to beta decreases and score downgrades are related to beta increases. Significant relationships are observed almost exclusively for the post-score-changes periods, supporting the hypothesis that extra-financial performance scores are leading indicators of systematic risk (*H1a*). Further, the leading effect is mostly associated with score downgrades, as opposed to upgrades (with the exception of score upgrades for the governance dimension), that are leading indicators of beta decreases. Specifically, we can observe that downgrades for almost all extra-financial scores are followed by significant increases in betas. For example, under the conditional model, governance performance score downgrades are followed by beta increases of 0.014 ($t = 2.38$), 0.021 ($t = 3.74$) and 0.017 ($t = 1.94$) for the [0; +60], [0; +120], and [0; +250] periods following score changes, respectively. The pattern is similar for downgrades in the social and environmental dimensions, as well as for the aggregate score. These results are consistent with an asymmetrical

relationship between extra-financial performance and systematic risk (*H2*).

Oikonomou *et al* (2012) examine the association between corporate social performance and financial risk for S&P 500 companies between the years 1992 and 2009 and also find that CSR is negatively but weakly related to systematic firm risk and that corporate social *irresponsibility* is positively and strongly related to financial risk. Their results, as well as ours, are consistent with the stylized fact according to which financial markets react more strongly to bad news. For example, De Bondt and Thaler (1985) show that investors "overreact" to unexpected and dramatic news events. Avouyi-Dovi and Neto (2004) observe that the asymmetric reaction to the signs of shocks can be explained by market participants' long positions on equity markets that would make them more sensitive to negative shocks.

The observed leading effect is the strongest for downgrades in the environment dimension. For example, under the conditional model, environment score downgrades are related to beta increases of 0.049 ($t = 7.09$) for the 120-day period following score downgrades, and related to increases of 0.016 ($t = 1.93$), 0.021 ($t = 3.74$) and 0.011 ($t = 1.90$) when downgrades involve social, governance, and aggregate scores, respectively. Increasing betas following extra-financial performance score downgrades (i.e., negative stakeholder information) is consistent with the view that there is a substantial and non negligible cost associated with environmental irresponsibility.

Overall, we find that beta variations following extra-financial performance score changes are economically and statistically significant. These results validate hypothesis *H1a*, which postulates that extra-financial performance scores are leading indicators of systematic risk. Further, the leading effect is asymmetrical, as most significant relationships are observed following score downgrades, which is consistent with hypothesis *H2*.

Table 3: Impact of extra-financial performance changes on systematic risk

Analysis period	Score upgrades						Score downgrades									
	Unconditional model			Conditional model			Unconditional model			Conditional model						
	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.				
[-120; 0]	-0.002 (-0.29)	0.004 (0.13)	-0.001 (-0.00)	-0.001 (-0.32)	-0.003 (-0.21)	0.017 (0.48)	-0.002 (-0.26)	-0.042 (-0.09)	0.012 (0.48)	0.035 (0.91)	0.004 (0.11)	0.001 (0.06)	0.010 (0.65)	0.022 (1.06)	0.002 (1.01)	0.004 (0.35)
[-60; 0]	-0.006 (-0.15)	0.008 (0.63)	-0.001 (-0.07)	-0.003 (-0.01)	-0.004 (-0.46)	-0.003 (-0.33)	-0.002 (-0.17)	-0.013 (-0.38)	0.009 (0.82)	0.042 (1.59)	0.006 (0.34)	0.012 (1.89)	0.008 (1.11)	0.039 (1.47)	0.003 (1.04)	0.009 (1.72)
[0; +60]	-0.006 (-1.23)	0.001 (0.09)	-0.014 (-1.47)	-0.016 (-1.88)	-0.007 (-1.19)	-0.011 (-1.42)	-0.011 (-1.01)	-0.009 (-1.71)	0.013 (1.92)	0.038 (3.83)	0.002 (0.48)	0.018 (2.61)	0.010 (1.97)	0.035 (2.38)	0.006 (0.84)	0.014 (2.38)
[0; +120]	-0.004 (-0.26)	-0.023 (-1.60)	-0.018 (-1.50)	-0.015 (-4.71)	-0.005 (-1.55)	-0.002 (-0.26)	-0.021 (-1.38)	-0.013 (-4.11)	0.018 (2.33)	0.062 (8.41)	0.020 (2.74)	0.023 (4.11)	0.011 (1.90)	0.049 (7.09)	0.016 (1.93)	0.021 (3.74)
[0; +250]	-0.003 (-1.43)	-0.005 (-1.11)	-0.028 (-1.33)	-0.011 (-2.46)	-0.004 (-1.29)	-0.002 (-1.14)	-0.023 (-1.33)	-0.009 (-2.58)	0.011 (2.49)	0.041 (5.92)	0.027 (5.62)	0.015 (1.82)	0.008 (1.72)	0.036 (4.29)	0.023 (5.02)	0.017 (1.94)
Obs.	1312	1104	930	867	1312	1104	930	867	901	772	521	860	901	772	521	860

Note: This table presents the results for the estimation of model (2). Mean values for b_{E1} are shown, with t-statistics in parentheses. Other coefficients in model (2) are estimated but not shown to save valuable space. Model (2) is estimated on a subsample of score upgrades and a subsample of score downgrades and, in each case, for the four types of score dimensions [aggregate (agg.), environment (envir.), social, and governance (gov.)]. Five analysis periods are considered ([-120; 0], [-60; 0], [0; +60], [0; +120], and [0; +250]), for which the estimation periods are, respectively: [-500; -120], [-500; -60], [-500; +60], [-500; +120], and [-500; +250]. The overall sample includes 266 Canadian firms from January 2007 to December 2012. Numbers in bold indicate significance at the 10 per cent level.



Conditioning on the current extra-financial performance of the firm

One of the problems with focusing on average beta variations estimated over the full sample is that it can conceal cases or special circumstances for which extra-financial score upgrades (i.e., positive stakeholder information) are related to beta increases. For instance, as previously outlined, we could observe that, dependent on certain conditions (e.g., overinvestment), CSR projects can increase systematic risk. To explore this conditionality further, we separate the firms in our sample into quintiles according to their extra-financial scores. We then estimate model (2) on a subsample of highly scored firms, defined as firms in the highest quintile in terms of extra-financial performance, and on a subsample of lowly scored firms, defined as firms in the lowest quintile in terms of extra-financial performance.

Table 4 provides results for highly scored firms, which show that extra-financial score upgrades for firms with already-high score predict higher systematic risk, particularly for the environmental and social dimensions, as well as for the aggregate score. In contrast, score downgrades have no significant impact on beta, except for the governance dimension for which score downgrades are related to beta increases in the 250-day period following score changes. These results support hypothesis *H3* and suggest that, for firms with already-high extra-financial performance, further improvements can be counterproductive and lead to an increase in systematic risk, possibly because of costs that investors feel are too high and inopportune. Our results also suggest that extra-financial irresponsibility (with the exception of the governance dimension) is not related to any future systematic risk increase if the firm already has a high extra-financial score.

Our results corroborate those from previous studies, including McWilliams and Siegel (2001) who determine that an optimal level of social performance exists beyond which it is less likely to shield the firm against the uncertainty and vulnerability of future cash flows. At extremely high levels of social performance, the drawbacks (costs) of CSR programs may outweigh the advantages (Handelman and Arnold, 1999; Smith, 2003), eventually leading to an increase in systematic risk for those firms. Other studies take an agency-cost perspective and express a negative view on the managerial motivations for pursuing CSR (e.g., Jensen and Meckling, 1976; Friedman, 1970; McWilliams *et al*, 2006; Krüger, 2015). These studies argue that managers may opportunistically use CSR to advance their careers or other personal agenda. Hemingway and Maclagan (2004) argue that one motivation for companies to adopt CSR is to cover up corporate misbehavior. The infamous firm *Enron*, for example, was widely viewed as a model of CSR and won several national awards for its environmental and community programs while at the same time engaging in massive accounting frauds that lead to its collapse in 2001 (Bradley, 2009). If firms use CSR as a tool to disguise bad news and divert shareholder scrutiny, CSR may then be associated with higher financial risk.

By contrast, according to Table 5, which provides results for firms with low extra-financial performance, we see that score upgrades are associated with lower systematic risk while score downgrades predict higher systematic risk. The leading effect of score upgrades for low-score firms is also greater in terms of magnitude and significance than for the full sample of firms. These results suggest that the efforts of low-score firms to improve their social image are related to future decreases in their systematic risk, while their social irresponsibility is related to increases in their systematic risk.

Table 4: Beta changes around changes in extra-financial performance for highly scored firms

Analysis period	Score upgrades						Score downgrades									
	Unconditional model			Conditional model			Unconditional model			Conditional model						
	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.				
[-120; 0]	0.008 (0.93)	0.022 (0.33)	0.036 (0.25)	-0.012 (-1.36)	0.002 (0.68)	0.017 (0.68)	0.030 (1.02)	-0.010 (-1.24)	-0.011 (-0.71)	-0.001 (-0.00)	-0.053 (-0.21)	0.024 (0.40)	0.000 (-0.00)	-0.003 (-0.01)	-0.046 (-0.05)	0.020 (0.24)
[-60; 0]	0.003 (0.71)	0.031 (1.11)	0.054 (1.02)	-0.022 (-1.62)	0.004 (0.24)	0.039 (1.06)	0.042 (0.55)	-0.022 (-1.27)	-0.009 (-0.35)	-0.002 (-0.55)	-0.084 (-0.20)	0.029 (1.01)	-0.006 (-0.16)	-0.001 (-0.41)	-0.070 (-0.02)	0.023 (0.58)
[0; +60]	0.017 (1.32)	0.036 (1.08)	0.139 (8.65)	-0.023 (-1.62)	0.016 (1.51)	0.022 (1.05)	0.133 (8.02)	-0.021 (-1.32)	-0.004 (-0.12)	-0.008 (-0.37)	-0.078 (-0.99)	0.055 (0.37)	-0.001 (-0.04)	-0.005 (-0.31)	-0.063 (-0.79)	0.066 (0.41)
[0; +120]	0.023 (2.53)	0.044 (3.72)	0.176 (13.70)	-0.014 (-1.64)	0.020 (2.29)	0.043 (3.48)	0.165 (12.59)	-0.011 (-1.40)	-0.005 (-0.40)	-0.009 (-0.56)	-0.062 (-0.80)	0.066 (1.64)	-0.008 (-0.69)	-0.006 (-0.06)	-0.068 (-0.63)	0.062 (1.17)
[0; +250]	0.019 (1.96)	0.030 (2.98)	0.130 (9.95)	-0.017 (-1.39)	0.021 (2.01)	0.029 (2.82)	0.127 (9.53)	-0.016 (-1.07)	-0.034 (-0.22)	-0.005 (-0.34)	-0.054 (-0.11)	0.046 (4.72)	-0.024 (-0.19)	-0.001 (-0.17)	-0.041 (-0.71)	0.045 (4.63)
Obs.	261	220	185	173	261	220	185	173	180	154	103	171	180	154	103	171

Note: This table presents the results for the estimation of model (2) on a subsample of highly scored firms. Highly-score firms are defined as firms in the highest quintile in terms of extra-financial score. Mean values for b_{Ei} are shown, with t statistics in parentheses. Other coefficients in model (2) are estimated but not shown to save valuable space. Model (2) is estimated on a subsample of score upgrades and a subsample of score downgrades and, in each case, for the four types of score dimensions [aggregate (agg.), environment (envir.), social, and governance (gov.)]. Five analysis periods are considered ([-120; 0], [-60; 0], [0; +60], [0; +120], and [0; +250]), for which the estimation periods are, respectively: [-500; -120], [-500; -60], [-500; +60], [-500; +120], and [-500; +250]. The overall sample includes 266 Canadian firms from January 2007 to December 2012. Numbers in bold indicate significance at the 10 per cent level.



Table 5: Beta changes around changes in extra-financial performance for firms with low extra-financial scores

Analysis period	Score upgrades						Score downgrades									
	Unconditional model			Conditional model			Unconditional model			Conditional model						
	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.				
[-120; 0]	0,002 (0,14)	0,000 (0,00)	0,001 (0,02)	-0,002 (-0,22)	0,003 (0,23)	0,002 (0,07)	0,002 (0,09)	-0,001 (-0,04)	0,002 (0,13)	0,005 (1,07)	0,001 (0,05)	0,003 (0,12)	0,001 (0,06)	0,002 (0,01)	0,002 (0,03)	
[-60; 0]	-0,013 (-0,15)	-0,034 (-1,12)	-0,004 (-0,01)	-0,003 (-0,01)	-0,020 (-0,99)	-0,029 (-1,11)	-0,002 (-0,17)	-0,001 (-0,01)	0,011 (1,01)	0,035 (1,03)	0,011 (0,78)	0,018 (1,69)	0,010 (0,33)	0,029 (1,03)	0,008 (0,30)	0,012 (1,61)
[0; +60]	-0,009 (-2,07)	-0,030 (-1,18)	-0,018 (-2,11)	-0,019 (-2,77)	-0,009 (-1,95)	-0,023 (-1,51)	-0,016 (-1,91)	-0,013 (-2,60)	0,019 (2,34)	0,041 (4,28)	0,020 (1,94)	0,022 (3,04)	0,015 (2,62)	0,039 (3,02)	0,014 (1,81)	0,020 (2,77)
[0; +120]	-0,008 (-2,76)	-0,027 (-2,41)	-0,020 (-2,74)	-0,016 (-4,98)	-0,006 (-2,29)	-0,021 (-2,09)	-0,017 (-2,41)	-0,014 (-4,52)	0,021 (2,99)	0,065 (7,32)	0,023 (2,74)	0,026 (4,89)	0,013 (2,03)	0,054 (7,11)	0,019 (2,78)	0,023 (3,90)
[0; +250]	-0,006 (-2,23)	-0,021 (-1,91)	-0,019 (-2,41)	-0,014 (-2,68)	-0,005 (-1,97)	-0,019 (-1,85)	-0,015 (-2,27)	-0,011 (-2,61)	0,015 (3,14)	0,045 (6,03)	0,032 (6,18)	0,021 (2,28)	0,011 (2,65)	0,039 (5,11)	0,027 (5,82)	0,019 (2,04)
Obs.	263	222	187	175	263	222	187	175	181	156	105	169	181	156	105	169

Note: This table presents the results for the estimation of model (2) on a subsample of firms with low extra-financial scores. Lowly scored firms are defined as firms in the lowest quintile in terms of extra-financial score. Mean values for b_{EI} are shown, with t-statistics in parentheses. Other coefficients in model (2) are estimated but not shown to save valuable space. Model (2) is estimated on a subsample of score upgrades and a subsample of score downgrades and, in each case, for the four types of score dimensions [aggregate (agg.), environment (envir.), social, and governance (gov.)]. Five analysis periods are considered ([-120; 0], [-60; 0], [0; +60], [0; +120], and [0; +250]), for which the estimation periods are, respectively, [-500; -120], [-500; -60], [-500; +60], [-500; +120], and [-500; +250]. The overall sample includes 266 Canadian firms from January 2007 to December 2012. Numbers in bold indicate significance at the 10 per cent level.

Robustness tests

Three robustness tests are conducted to ensure the validity of our results in different circumstances. The first test examines the impact of extra-financial performance changes on risk on a subsample of firms that are not contaminated by any firm-specific event. The second test examines the potentially different impact of extra-financial performance score upgrades and downgrades on risk according to the nature of the firm's business. The third test investigates the potentially different impact of extra-financial performance score upgrades and downgrades on risk depending on market conditions.

Impact of extra-financial performance score changes on risk for a subsample of uncontaminated firms

Long-term studies are sensitive to the presence of confounding effects because other value-relevant events, which are not necessarily related to CSR, can occur throughout a given year. To ensure that our results are not attributable to other firm-specific events, we re-estimate model (2) on a subsample of firms for which no important event takes place during the analysis period. To do so, we use Bloomberg data to estimate differences between quarterly earnings announcements for firms in the TSX and market expectations prior to these announcements. Following the literature (e.g., Mendenhall, 2004), we normalize these differences by their standard deviation over the period under study and select events with the largest absolute value as proxies for potentially contaminating events. This approach removes approximately 15 per cent of our observations. Results for the estimation of model (2) on the remaining uncontaminated observations are available in Table 6 and are very similar, even stronger, than results obtained previously.

Impact of extra-financial performance score upgrades and downgrades on risk conditional on the firm's business sector

In this section, we consider the possibility that the relationship between extra-financial performance and systematic risk is heterogeneous across industries.⁹ Specifically, there is empirical evidence suggesting that firm risk varies by industry (Fama and French, 1997; Gebhardt *et al*, 2001). In addition, some studies show that extra-financial performance varies significantly across sectors (e.g., Carroll, 1979; Griffin and Mahon, 1997; Brammer *et al*, 2006; Godfrey *et al*, 2008). To verify if the impact of extra-financial performance depends on the business sector of the firm, we examine the impacts of score upgrades and downgrades on systematic risk for each of the ten industrial sectors in our sample, based on the Global Industry Classification Standard (GICS) of each firm. Specifically, we estimate model (2) separately for each of the ten industries. Table 7 reports the estimates for b_{Ei} for each industry for the [0; +250] period. Results for are available in Table 7. We first note that score downgrades are strongly related to increases in systematic risk (beta) for all industries. Secondly, there is no clear evidence that the impact of extra-financial performance differs according to the firm's business sector. For instance, for the aggregate score, conditional model coefficients range from 0.011 to 0.027. We nevertheless note that the effect of environmental-dimension score downgrades on beta is the highest for the following industries: (i) Energy (conditional model coefficient of 0.059) which includes the oil and gas sectors; (ii) Materials (0.057) which include the metals and mining sector, and (iii) Industrials (0.055) which include the airline, marine, road, and rail sector. The fact that these industries are particularly exposed to environmental issues may explain the relatively stronger effect for the environmental dimension of extra-financial performance. In a similar matter, we observe that the effect of



Table 6: Changes in firm beta around changes in extra-financial performance scores on a subsample of uncontaminated firms

Analysis period	Score upgrades						Score downgrades									
	Unconditional model			Conditional model			Unconditional model			Conditional model						
	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.				
[−120; 0]	−0.007 (−0.83)	−0.002 (−0.47)	−0.002 (−0.11)	−0.009 (−0.25)	−0.004 (−0.21)	−0.001 (−0.83)	−0.003 (−0.26)	−0.002 (−0.03)	0.017 (0.73)	0.036 (1.01)	0.008 (0.14)	0.002 (0.01)	0.021 (0.99)	0.025 (1.19)	0.003 (1.13)	0.001 (0.04)
[−60; 0]	−0.001 (−0.99)	−0.008 (−0.89)	−0.003 (−0.07)	−0.009 (−0.14)	−0.008 (−0.73)	−0.003 (−0.93)	−0.002 (−0.28)	−0.011 (−0.67)	0.005 (0.73)	0.047 (1.61)	0.004 (0.22)	0.004 (1.19)	0.011 (1.29)	0.042 (1.59)	0.093 (1.20)	0.003 (1.04)
[0; +60]	−0.006 (−1.37)	−0.002 (−1.22)	−0.023 (−1.49)	−0.011 (−1.99)	−0.007 (−1.31)	−0.015 (−1.26)	−0.019 (−1.36)	−0.012 (−2.23)	0.014 (2.48)	0.041 (4.07)	0.008 (0.79)	0.021 (3.14)	0.013 (3.02)	0.037 (2.73)	0.005 (0.24)	0.017 (2.97)
[0; +120]	−0.044 (−1.32)	−0.030 (−1.60)	−0.022 (−1.61)	−0.018 (−5.02)	−0.005 (−1.43)	−0.029 (−1.21)	−0.021 (−1.42)	−0.015 (−4.87)	0.019 (2.76)	0.065 (8.66)	0.022 (2.90)	0.026 (4.69)	0.013 (2.33)	0.053 (7.73)	0.018 (2.56)	0.023 (4.03)
[0; +250]	−0.004 (−1.53)	−0.027 (−1.29)	−0.031 (−1.39)	−0.014 (−2.93)	−0.004 (−1.36)	−0.027 (−0.89)	−0.033 (−1.36)	−0.011 (−2.81)	0.015 (2.81)	0.044 (6.03)	0.029 (6.13)	0.019 (2.63)	0.011 (2.21)	0.039 (4.93)	0.023 (5.41)	0.018 (2.16)
Obs.	1105	970	866	783	1105	970	866	783	769	688	483	807	769	688	483	807

Note: This table presents the results for the estimation of model (2) on a subsample of firms for which there are no surprises in quarterly earnings announcements during the analysis period. Surprises are defined as differences between quarterly earnings announcements for firms in the TSX and market expectations prior to these announcements as estimated by Bloomberg. Mean values for b_{Ei} are shown, with t statistics in parentheses. Other coefficients in model (2) are estimated but not shown to save valuable space. Model (2) is estimated on a subsample of score downgrades and a subsample of score upgrades and, in each case, for the four types of score dimensions [aggregate (agg.), environment (envir.), social, and governance (gov.)]. Five analysis periods are considered ([−120; 0], [−60; 0], [0; +60], [0; +120], and [0; +250]), for which the estimation periods are, respectively, [−500; −120], [−500; −60], [−500; +60], [−500; +120], and [−500; +250]. The overall sample includes 266 Canadian firms from January 2007 to December 2012. Numbers in bold indicate significance at the 10 per cent level.

Table 7: Changes in firm beta around changes in extra-financial performance scores conditional on the business sector

Sector	Score upgrades						Score downgrades									
	Unconditional model			Conditional model			Unconditional model			Conditional model						
	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.				
Consumer discretionary	-0.006 (-1.53)	-0.002 (-0.11)	-0.028 (-2.09)	-0.005 (-1.26)	-0.004 (-1.39)	-0.001 (-0.05)	-0.025 (-1.98)	-0.003 (-0.97)	0.016 (3.21)	0.040 (3.91)	0.037 (6.22)	0.011 (3.02)	0.015 (2.98)	0.037 (3.42)	0.035 (6.03)	0.010 (2.96)
Consumer staple	-0.005 (-1.46)	-0.001 (-0.07)	-0.029 (-2.33)	-0.003 (-1.02)	-0.004 (-1.30)	-0.001 (-0.03)	-0.028 (-2.03)	-0.002 (-0.61)	0.017 (3.50)	0.041 (3.96)	0.039 (6.53)	0.010 (3.29)	0.016 (3.39)	0.039 (3.59)	0.038 (6.23)	0.009 (2.95)
Energy	-0.008 (-1.66)	-0.015 (-1.79)	-0.005 (-1.13)	-0.012 (-3.19)	-0.006 (-1.49)	-0.012 (-1.62)	-0.004 (-1.02)	-0.010 (-3.01)	0.029 (5.22)	0.062 (6.93)	0.030 (5.69)	0.017 (5.81)	0.027 (5.01)	0.059 (6.23)	0.029 (5.32)	0.015 (5.17)
Financials	-0.007 (-1.68)	-0.003 (-0.11)	-0.019 (-1.68)	-0.017 (-4.63)	-0.006 (-1.37)	-0.002 (-0.09)	-0.017 (-1.53)	-0.015 (-4.12)	0.023 (3.94)	0.037 (2.92)	0.027 (3.91)	0.020 (6.83)	0.021 (3.63)	0.036 (2.29)	0.024 (3.72)	0.018 (6.14)
Health care	-0.001 (-0.13)	-0.010 (-1.69)	-0.003 (-1.04)	-0.005 (-1.14)	-0.001 (-0.09)	-0.008 (-1.62)	-0.002 (-0.96)	-0.003 (-1.01)	0.013 (2.19)	0.043 (5.97)	0.025 (3.71)	0.010 (3.15)	0.011 (2.09)	0.041 (3.78)	0.023 (3.53)	0.009 (2.93)
Industrials	-0.005 (-1.51)	-0.010 (-1.77)	-0.008 (-1.63)	-0.011 (-2.10)	-0.004 (-1.41)	-0.008 (-1.60)	-0.006 (-1.51)	-0.009 (-1.99)	0.019 (3.44)	0.057 (6.73)	0.026 (3.82)	0.015 (4.73)	0.018 (3.39)	0.055 (5.61)	0.025 (3.35)	0.014 (4.12)
Information technology	-0.003 (-1.03)	-0.004 (-1.09)	-0.006 (-1.11)	-0.002 (-0.66)	-0.002 (-0.23)	-0.003 (-0.95)	-0.004 (-1.09)	-0.001 (-0.31)	0.013 (2.47)	0.038 (3.35)	0.023 (3.22)	0.011 (3.07)	0.012 (2.16)	0.036 (3.09)	0.021 (3.10)	0.009 (2.89)
Materials	-0.006 (-1.61)	-0.014 (-1.83)	-0.009 (-1.41)	-0.012 (-2.67)	-0.005 (-1.29)	-0.011 (-1.61)	-0.006 (-1.53)	-0.011 (-2.13)	0.022 (3.74)	0.058 (6.83)	0.031 (5.82)	0.016 (4.96)	0.020 (3.59)	0.057 (5.69)	0.030 (5.65)	0.015 (4.20)
Telecommunication services	-0.003 (-1.00)	-0.003 (-0.89)	-0.003 (-1.03)	-0.005 (-1.09)	-0.002 (-0.29)	-0.002 (-0.10)	-0.002 (-0.93)	-0.003 (-0.78)	0.015 (2.52)	0.036 (3.14)	0.027 (3.89)	0.009 (2.92)	0.012 (2.41)	0.036 (2.97)	0.025 (3.74)	0.008 (2.63)
Utilities	-0.003 (-1.08)	-0.005 (-1.13)	-0.005 (-1.10)	-0.003 (-0.40)	-0.003 (-1.06)	-0.003 (-1.11)	-0.004 (-1.00)	-0.001 (-0.13)	0.014 (2.33)	0.035 (2.97)	0.026 (3.81)	0.009 (2.81)	0.011 (2.06)	0.034 (2.37)	0.024 (3.69)	0.007 (2.63)

Note: This table presents the results for the estimation of model (2) conditional on the business sector of the firm. Mean values for b_{EI} are shown, with t statistics in parentheses. Other coefficients in model (2) are estimated but not shown to save valuable space. Model (2) is estimated on a subsample of score upgrades and a subsample of score downgrades and, in each case, for the four types of score dimensions [aggregate (agg.), environment (envir.), social, and governance (gov.)]. Five analysis periods are considered ([-120; 0], [-60; 0], [0; +60], [0; +120], and [0; +250]), for which the estimation periods are, respectively: [-500; -120], [-500; -60], [-500; +60], [-500; +120], and [-500; +250]. The overall sample includes 266 Canadian firms from January 2007 to December 2012. Numbers in bold indicate significance at the 10 per cent level.



Table 8: Changes in firm beta around changes in extra-financial performance scores conditional on market conditions

Analysis period	Score upgrades (Conditional model)						Score downgrades (Conditional model)									
	2007–2009			2010–2012			2007–2009			2010–2012						
	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.	Agg.	Envir.	Social	Gov.				
[–120; 0]	–0.004 (–0.98)	–0.001 (–0.11)	–0.001 (–0.13)	–0.005 (–0.03)	–0.003 (0.48)	0.003 (0.48)	0.000 (–0.01)	–0.004 (–0.14)	0.009 (0.43)	0.027 (1.11)	0.004 (1.08)	0.002 (0.01)	0.010 (0.99)	0.023 (1.00)	0.001 (0.52)	0.001 (0.89)
[–60; 0]	–0.006 (–1.23)	–0.003 (–0.74)	–0.004 (–0.77)	–0.012 (–1.50)	–0.004 (–0.29)	–0.004 (–0.14)	–0.001 (–0.22)	–0.009 (–1.22)	0.008 (1.45)	0.040 (1.63)	0.008 (1.13)	0.018 (2.66)	0.008 (1.59)	0.039 (1.35)	0.002 (0.98)	0.007 (1.04)
[0; +60]	–0.008 (–1.47)	–0.008 (–1.51)	–0.013 (–1.24)	–0.013 (–3.68)	–0.006 (–1.33)	–0.010 (–1.39)	–0.011 (–1.30)	–0.012 (–1.90)	0.011 (2.97)	0.037 (3.07)	0.0015 (1.53)	0.025 (4.55)	0.010 (2.03)	0.035 (2.44)	0.008 (0.83)	0.012 (2.63)
[0; +120]	–0.006 (–1.67)	–0.005 (–1.27)	–0.024 (–1.41)	–0.016 (–4.11)	–0.005 (–1.27)	–0.004 (–1.03)	–0.022 (–1.21)	–0.014 (–2.19)	0.012 (2.83)	0.050 (7.74)	0.018 (2.49)	0.029 (4.98)	0.011 (2.58)	0.049 (7.17)	0.016 (2.06)	0.011 (2.37)
[0; +250]	–0.008 (–1.63)	–0.004 (–1.52)	–0.025 (–1.63)	–0.012 (–4.86)	–0.006 (–1.01)	–0.003 (–1.23)	–0.020 (–1.06)	–0.010 (–2.97)	0.009 (2.89)	0.037 (4.89)	0.025 (5.49)	0.024 (3.28)	0.008 (2.43)	0.035 (4.08)	0.022 (4.96)	0.010 (2.01)
Obs.	583	282	291	401	729	822	639	466	340	156	157	255	561	616	364	605

Note: This table presents the results for the estimation of model (2) conditional on market conditions. The sample is split into two subperiods: 2007–2009 and 2010–2012 and model (2) is estimated for each subperiod. Mean values for b_{EI} are shown, with t statistics in parentheses. Other coefficients in model (2) are estimated but not shown to save valuable space. Model (2) is estimated on a subsample of score upgrades and a subsample of score downgrades and, in each case, for the four types of score dimensions [aggregate (agg.), environment (envir.), social, and governance (gov.)]. Five analysis periods are considered ([–120; 0], [–60; 0], [0; +60], [0; +120], and [0; +250]), for which the estimation periods are, respectively, [–500; –120], [–500; –60], [–500; +60], [–500; +120], and [–500; +250]. The overall sample includes 266 Canadian firms from January 2007 to December 2012. Numbers in bold indicate significance at the 10 per cent level.



social-dimension score downgrades seems slightly stronger for industries related to retail and light manufacturing such as consumer discretionary (conditional model coefficient of 0.035) and consumer staple (0.038). One potential explanation is that employee relations and human rights issues are very important for these labor-intensive industries.

In terms of score upgrades, there is weak evidence, mostly for the unconditional model, that score upgrades are associated with beta changes. However, some industries do appear to be more affected than others by extra-financial performance changes. For instance, only the two retail industries, namely consumer discretionary and consumer staple, experience a significant reduction in systematic risk following social-dimension score upgrades (for the conditional model). Similarly, only four industries are affected by governance-dimension extra-financial performance changes: financials, energy, industrials and materials. In the four cases, systemic risk is lower following score upgrades.

Impact of extra-financial performance score upgrades and downgrades on risk conditional on the economic context

With our third robustness test, we wish to investigate whether market conditions can mitigate or amplify the impact of extra-financial performance score upgrades and downgrades on risk. Recent studies (e.g., Albuquerque *et al*, 2014; Oikonomou, 2012; Chen *et al*, 2010) show that extra-financial performance and macroeconomic factors can co-vary.¹⁰ Further, our earlier results suggest that systematic risk is a function of the economic context. It is therefore possible that extra-financial performance changes have a different impact on risk depending on market conditions. This analysis is particularly interesting since the period covered by our study, from 2007 to 2012, is marked by the subprime financial crisis.

To test whether the impact of extra-financial performance score changes on risk depends on

market conditions, we split our sample into two subperiods: (i) the 2007–2009 subperiod, which corresponds to the financial crisis, and (ii) the 2010–2012 subperiod, which corresponds to the post-crisis relatively stable period.¹¹ We estimate model (2) for each subperiod. Results in Table 8 show that the estimates of coefficient b_{Ei} are remarkably stable over time, suggesting that there is very little significant evidence that the impact of extra-financial performance varies according to market conditions. We nevertheless note that the effect of governance score changes, especially score downgrades, are greater for the 2007–2009 subperiod than for the 2010–2012 subperiod. This suggests that, during times of economic uncertainty, governance-dimension-downgraded firms experience higher systematic risk, but governance-dimension upgraded firms are not rewarded, at least in terms of systematic risk, by the market. Corporate governance therefore appears to be a greater concern for investors during a financial crisis.

CONCLUSION

This study investigates the informational content of extra-financial performance scores by examining the relationship between extra-financial score changes (upgrades and downgrades) and systematic risk (beta) variations. Our work is based on changes in Sustainalytics' extra-financial performance scores for a sample of 266 Canadian corporations between 2007 and 2012 and provides important empirical findings. First, we find no significant evidence that changes in extra-financial performance scores lag beta variations. Rather, we find that systematic risk increases follow extra-financial score downgrades. Extra-financial scores therefore do not appear to be established a posteriori on the basis of stock market information but rather appear to be leading indicators of systematic risk variations.

Our results also show that score upgrades for firms with already-high scores predict higher



systematic risk, while score upgrades predict lower systematic risk for lower-scored firms. This result suggests that, for firms with already-high extra-financial scores, improvements can be counterproductive and lead to an increase in systematic risk, possibly because of costs that investors feel are too high and inopportune. However, systematic risk decreases when firms with low extra-financial scores make an effort to improve their social image. Finally, while score downgrades are not related to beta variations for firms with high extra-financial scores, they are related to higher systematic risk for firms with low extra-financial scores.

Overall, this study provides evidence of the usefulness of extra-financial agencies' scorings for managers in the development of their risk management strategies. Investors may limit their exposure to systematic risk by following changes in firms' ESG ratings. They can also build investment strategies based on the changes in corporate extra-financial performance score. Our results also imply that it may be appropriate to include an irresponsibility risk factor in a general asset pricing model. We leave this question and tests to future research.

One of limitations of this study is that Sustainalytics from which the scores are collected, may be one of the largest providers of corporate social responsibility intelligence in Canada, but is nonetheless only one agency among many.

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NOTES

1. Formerly KLD Research & Analytics, Inc.
2. Ethical Investment Research and Information Service (<http://www.eiris.org/>).

3. ASSET4 provides investment research information on economic, environmental, social, and governance (ESG) aspects of corporate performance.
4. Sustainalytics was formed from the merger between the Dutch firm "Sustainalytics" and the Canadian firm "Jantzi Research Inc" in August 2009 (<http://www.sustainalytics.com/>).
5. We use the term "extra-financial" performance to include all types of non financial performance that are deemed "responsible". These include the social, governance and environmental performances of firms.
6. Christopherson *et al* (1998) show that a conditional approach, using time-varying measures of risk (beta) and abnormal performance (alpha), is better able to predict future performance than conditional beta models that consider only time-varying measures of risk (Ferson and Schadt, 1996).
7. We follow a two-step methodology to select the variables that reflect the Canadian economy. In the first step, based on the literature (e.g., Ferson and Qian, 2004), which mainly focuses on the U.S. economy, we identify seven macrofinancial information variables intended to reflect the state of the economy: (i) short-term interest rates, (ii) interest rate volatility, (iii) term structure of interest rates, (iv) term structure concavity, (v) stock market performance, (vi) stock market (implied) volatility, and (vii) credit spread. In the second step, we use stepwise regression techniques to identify those information variables that have a predictive power on financial performance and beta. In the end, we identify five information variables, $Z_{n,t}$, that are used throughout: (i) short-term interest rates, (ii) term structure of interest rates, (iii) stock market performance, (iv) stock market (implied) volatility, and (v) credit spread.
8. The other parameters in model (2) are estimated and included in the regressions, but are not reported.
9. The authors would like to thank an anonymous referee for this suggestion.
10. Given the cost of fulfilling ESG criteria during difficult economic times, firms may reduce their investments in CSR initiatives (Albuquerque *et al*, 2014) and thereby appear to be less observant of CSR criteria which may negatively affect their extra-financial performance score. On the other hand, some authors (see, for e.g., Oikonomou *et al*, 2012) believe that a higher extra-financial performance score should be expected during periods of economic uncertainty when firms may be more inclined to implement good practices, including socially responsible ones, to reduce risk (Chen *et al*, 2010).
11. As in Aloui *et al* (2011), we choose the subperiod of 2007–2009 as representative of the financial subprime crisis. As highlighted by Longstaff (2010), the subprime crisis actually began in early 2007. Moreover, this period includes the period of contraction from December 2007 to June 2009, as identified by NBER.

REFERENCES

- Aloui, R., Aïssa, M.S.B., Nguyen, D.K. (2011) Global financial crisis, extreme interdependences, and contagion effects: The role of economic structure? *Journal of Banking & Finance* 35: 130–141.



- Albuquerque, R.A., Durnev, A. and Koskinen, K. (2014) Corporate social responsibility and firm risk: Theory and empirical evidence. SSRN working paper.
- Avouyi-Dovi, S.D. and Neto, D. (2004) Equity market interdependence: the relationship between European and US stock markets. *Banque de France, Financial Stability Review* 4: 108–126.
- Barnea, A. and Rubin, A. (2010) Corporate social responsibility as a conflict between shareholders. *Journal of Business Ethics* 97: 71–86.
- Baron, D. and Diermeier, D. (2007) Strategic activism and nonmarket strategy. *Journal of Economics and Management Strategy* 16: 599–634.
- Bouslah, K., Kryzanowski, L. and M'Zali, B. (2013) The impact of the dimensions of social performance on firm risk. *Journal of Banking & Finance* 37: 1258–1273.
- Boutin-Dufresne, F. and Savaria, P. (2004) Corporate social responsibility and financial risk. *Journal of Investing* 13: 57–66.
- Bradley Jr., R.L. (2009) Corporate social responsibility and energy. In Horowitz, I.L. (Ed.) *Culture and Civilization*. Transaction Publishers, New Jersey, 181–197.
- Brammer, S., Brooks, C., Pavelin, S. (2006) Corporate social performance and stock returns: UK evidence from disaggregate measures. *Financial Management* 35: 97–116.
- Carroll, A.B. (1979) Three-dimensional conceptual model of corporate performance. *The Academy of Management Review* 4: 497–505.
- Cespa, G. and Cestone, G. (2007) Corporate social responsibility and managerial entrenchment. *Journal of Economic Management and Strategy* 16: 741–771.
- Chen, C., Guo, W. and Tay, N. (2010) Are member firms of corporate groups less risky? *Financial Management* 39: 59–82.
- Christopherson, J., Ferson, W. Glassman, D. (1998). Conditioning manager alphas on economic information: another look at the persistence of performance. *Review of Financial Studies* 11: 111–142.
- Champagne C., Chrétien S. and Coggins F. (2015) Effects of pension fund freezing on firm performance and risk. *Canadian Journal of Administrative Sciences*.
- De Bondt W.-M. and Thaler R. (1985) Does the stock market overreact? *Journal of Finance* 40: 793–805.
- Fama, E. and French, K. (1993) Common risk factors in the returns on stocks and bonds. *Journal of Financial Economics* 33: 3–56.
- Fama, E.F., French, K.R. (1997) Industry cost of equity. *Journal of Financial Economics* 43: 153–193.
- Ferson, W.E. and Qian, M. (2004) *Conditional Performance Evaluation Revisited*, Research Foundation Monograph of the CFA Institute, ISBN 0-943205-69-7.
- Ferson, W. and Schadt, R. (1996) Measuring fund strategy and performance in changing economic conditions. *Journal of Finance* 51: 425–462.
- Friedman, M., (1970) The social responsibility of business is to increase its profits. *New York Times*, September, 13, 122–126.
- Gebhardt, W., Lee, C., Swaminathan, B. (2001) Toward an implied cost of capital. *Journal of Accounting Research* 39: 135–176.
- Godfrey, P.C., Hatch, N.W., Hansen, J.M. (2008) Toward a general theory of CSRs: The roles of beneficence, profitability, insurance, and industry heterogeneity. *Business and Society*, 49: 316–344.
- Griffin, J.J., Mahon, J.F. (1997) The corporate social performance and corporate financial performance debate: Twenty-five years of incomparable research. *Business and Society* 36: 5–31.
- Handelman, J.M. and Arnold S.J. (1999) The role of marketing actions with a social dimension: Appeals to the institutional environment. *Journal of Marketing* 63: 33–48.
- Hasseldine, J., Salama, A. and Toms, J.S. (2005) Quantity versus quality: the impact of environmental disclosures on the reputations of UK Plcs. *British Accounting Review* 2: 231–248.
- Hemingway, C.A. and MacLagan, P.W. (2004) Managers' personal values as drivers of corporate social responsibility. *Journal of Business Ethics* 50: 33–44.
- Holthausen R.W. and Leftwic R.W. (1986) The effect of bond rating changes on common stock prices. *Journal of Financial Economics* 17(1): 57–89.
- Hong, H.G., Kubik, J.K., Scheinkman, J.A. (2012) Financial constraints on corporate goodness. Unpublished working paper. Princeton University and Syracuse University.
- Jensen, M., Meckling, W. (1976) Theory of the firm: Managerial behavior, agency cost and capital structure. *Journal of Financial Economics* 3: 305–360.
- Jo, H. and Na, H. (2012) Does CSR reduce firm risk? Evidence from controversial industry sectors. *Journal of Business Ethics* 110: 441–457.
- Jones, T.M. (1995) Instrumental stakeholder theory: A synthesis of ethics and economics'. *Academy of Management Review* 2: 404–437.
- Jorion, P., Zhang, G. (2007) Information effects of bond rating changes: The role of the rating prior to the announcement. *The Journal of Fixed Income* 16(4): 45–59.
- Kim, Y., Li, H. and Li, S. (2014) Corporate social responsibility and stock price crash risk. *Journal of Banking & Finance* 43: 1–13.
- Krüger, P. (2015) Corporate goodness and shareholder wealth. *Journal of Financial Economics* 115: 304–329.
- Lee, D.D. and Faff, R.W. (2009) Corporate sustainability performance and idiosyncratic risk: A global perspective. *Financial Review* 2: 213–237.
- Longstaff, F.A. (2010) The subprime credit crisis and contagion in financial markets. *Journal of Financial Economics* 97: 436–450.
- Luo, X. and Bhattacharya, C.B. (2009) The debate over doing good: Corporate social performance, strategic marketing levers, and firm-idiosyncratic risk. *Journal of Marketing* 73: 198–213.
- Lyon, T.P. and Maxwell, J.W. (2011) Greenwash: Corporate environmental disclosure under threat of audit. *Journal of Economics & Management Strategy* 20: 3–41.
- Mattingly, J.E. and Berman, S.L. (2006) Measurement of corporate social action: Discovering taxonomy in the Kinder Lydenburg domini ratings data. *Business & Society* 45: 20–46.
- McGuire, J.B., Schneeweis, T. and Sundgren, A. (1988) Corporate social responsibility and firm financial performance. *Academy of Management Journal* 3: 854–872.
- McWilliams, A. and Siegel, D. (2001) Corporate social responsibility: A theory of the firm perspective. *Academy of Management Review* 26: 117–127.



- McWilliams, A., Siegel, D., Wright, P. (2006) Guest editors' introduction corporate social responsibility: Strategic implications. *Journal of Management Studies* 43: 1–18.
- Mendenhall, R.R. (2004) Arbitrage risk and post-earnings-announcement drift. *The Journal of Business* 77: 875–894.
- Mishra, S. and Modi, S.B. (2012) Positive and negative corporate social responsibility, financial leverage, and idiosyncratic risk. *Journal of Business Ethics* 2: 431–448.
- Newey, W. and West, K.D. (1987) A simple, positive semi-definite heteroskedasticity and autocorrelation-consistent covariance matrix. *Econometrica* 55: 703–708.
- Oikonomou, I., Brooks, C. and Pavelin, S. (2012) The impact of corporate social performance on financial risk and utility: A longitudinal analysis. *Financial Management* 41: 483–515.
- Orlitzky, M. and Benjamin, J.D. (2001) Corporate social performance and firm risk: A meta-analytic review. *Business and Society* 4: 369–396.
- Porter, M.E. and Kramer, M.R. (2006) Strategy and society: the link between competitive advantage and corporate social responsibility. *Harvard Business Review* 12: 78–92.
- Preston, L.E. and O'Bannon, D.P. (1997) The corporate social-financial performance relationship. A typology and analysis. *Business and Society* 4: 109–125.
- Roberts, R.W. (1992) Determinants of corporate social responsibility disclosure: An application of stakeholder theory. *Accounting, Organizations and Society* 6: 595–612.
- Salama, A., Anderson, K.P. and Toms, S. (2011) Does community and environmental responsibility affect firm risk? Evidence from UK panel data 1994–2006. *Business Ethics* 20: 192–204.
- Sharfman, M.P. and Fernando, C.S. (2008) Environmental risk management and the cost of capital. *Strategic Management Journal* 6: 569–592.
- Smith, C.N. (2003) Corporate social responsibility: Whether or how? *California Management Review* 45: 52–76.
- Stern, N. (2006) *Stern Review. The Economics of Climate Change*. Cambridge: New Economics Foundation.
- Waddock, S.A. and Graves, S. B. (1997) The corporate social performance-financial performance link. *Strategic Management Journal* 18: 303–319.
- Weinstein, M.I. (1977) The effect of a rating change announcement on bond price. *Journal of Financial Economics* 5(3): 329–350.
- Wijnberg, N.M. (2000) Normative stakeholder theory and aristotle: the link between ethics and politics. *Journal of Business Ethics* 4: 329–342.

APPENDIX

See Table 9.



Table 9: Sustainability's ESG metrics

	Environment			Social		Governance
Operations	Formal environmental policy	Employees	Policy on freedom of association	Business ethics	Policy on bribery and corruption	
	Environmental management system		Formal policy on the elimination of discrimination		Signatory to UN global compact	
	External certification of EMS		Programs to increase workforce diversity		Tax transparency	
	Environmental fines and non-monetary sanctions		Percentage of employees covered by collective bargaining agreements		Business ethics related controversies or incidents	
	Participation in carbon disclosure project (CDF)		Employee turnover rate	Reporting, transparency and oversight	CSR reporting quality	
	Scope of corporate reporting on GHG emissions		Top employer recognition		External verification of CSR reporting	
	Programs and targets to reduce direct GHG emissions	Supply chain	Employee related controversies or incidents		Disclosure of directors' remuneration	
	Programs and targets to increase renewable energy use		Scope of social supply chain standards		Oversight of ESG issues	
	Carbon intensity		Contractors & supply chain related controversies or incidents		Executive compensation tied to ESG performance	
	Carbon intensity trend	Tenants	Public position statement on responsible marketing		Board diversity	
	per cent Primary energy use from renewables		Customer related controversies or incidents		Separation of board chair and CEO roles	
	Operations related controversies or incidents	Community and philanthropy	Activities in sensitive countries		Board independence	
Supply chain	Formal policy or programme on green procurement		Policy on human rights		Audit committee independence	
	contractors & supply chain related controversies or incidents		Society & community related controversies or incidents		Governance related controversies or incidents	
Products & services and sustainability	Sustainability related products & services		Guidelines for philanthropic activities and primary areas of support	Public policy	Transparency on payments to host governments	
	Products & services related controversies or incidents		Corporate foundation		Public policy related controversies or incidents	